

Learning Atom Probe Tomography time-of-flight peaks for mass-to-charge ratio spectrometry

Kevin J Coakley¹, Norman A Sanford

National Institute of Standards and Technology, 325 Broadway, Boulder CO 80305, USA

Abstract

In laser-assisted atom probe tomography, an important goal is to reconstruct the mass-to-charge ratio, (m/z) , spectrum due to various ion species. In general, the probability mass function (pmf) associated with the time-of-flight (TOF) spectrum produced by each ion species is unknown and varies from species-to-species. Moreover, measuring pmfs for distinct ion species in calibration experiments is not practical. Here, we present a mixture model method to determine TOF pmfs that can vary from peak-to-peak. In this approach, we determine weights of candidate pmfs with a maximum likelihood method. In a proof-of-principle study, we apply our method to a TOF spectrum acquired from a silicon sample and determine intensity estimates of singly charged isotopes of silicon.

Keywords: Atom Probe Tomography, Cross-validation, Expectation Maximization algorithm, Machine learning, Maximum likelihood, Mixture model, Silicon isotopes

1. Introduction

2 In laser-assisted atom probe tomography (APT) [1, 2, 3, 4, 5], a sharp,
3 needle-shaped specimen is subjected to an external electric field bias, which
4 is held just below the threshold for field evaporation of ions. Field evapora-
5 tion is triggered by incident laser pulses, imparting weak thermal transients,
6 enabling the evaporation threshold to be exceeded. As a result, surface atoms

Email addresses: kevin.coakley@nist.gov (Kevin J Coakley),
norman.sanford@nist.gov (Norman A Sanford)

¹Corresponding author

7 originating from the vicinity of the specimen apex are evaporated (in the form
 8 of ions) synchronously with the laser pulses. The ions hit a 2-dimensional de-
 9 tector and their respective impact locations and times-of-flight are recorded.
 10 A back-projection algorithm applied to the accumulated data enables 3-D
 11 reconstruction of the specimen. This reconstruction locates and identifies
 12 each constituent species. Given the mass m and charge z of the ion, in the
 13 simplest approximation, the theoretical time-of-flight (transit time) of the
 14 ion to the detector, t , satisfies the following equation

$$\frac{m}{z} = \frac{2eV_{app}}{L^2}t^2, \quad (1)$$

15 where e is the charge of the electron, V_{app} is the bias voltage between the
 16 specimen and the electrode and L is the length of the ion trajectory, which
 17 is referred to as the flight length.

18 In actual experiments, the specimens are often conical so that the apex
 19 diameter gradually increases as ions are evaporated away and the specimen
 20 shortens. In this context, “apex diameter” refers to the diameter of a sphere
 21 which approximates the shape of a specimen apex. As the apex diameter
 22 increases, V_{app} is generally increased to maintain a constant evaporation field
 23 for a desired detection rate of ions. For conical specimens, the initial and final
 24 apex diameters are typically in the range of 20 nm and 150 nm, respectively;
 25 the length reduction of the specimen during an APT session is often a few
 26 hundred nm depending upon experimental targets. The specimen-electrode
 27 separation is roughly 30 μm to 40 μm . Now, if V_{app} is instead held constant,
 28 there is a negligible reduction in the evaporation field due only to the sample
 29 shortening and the apex receding from the electrode as ions are evaporated
 30 away; the increasing apex diameter is the dominant effect that influences
 31 the evaporation field when V_{app} is held constant. However, for a rod-like
 32 specimen of uniform diameter, V_{app} and the apex diameter, can both remain
 33 nearly constant throughout the APT run while still maintaining the desired
 34 detection rate [7]. In such a situation, V_{app} will rise solely because the apex
 35 retreats from the electrode as ions are lost from the specimen. An analysis
 36 of APT specimen apex electrostatics has been given by Zhang et al. [8].

37 In our experiments, the electrode-to-detector separation is 90 mm. Within
 38 the approximations leading to Eq. 1, variations in flight length due to the
 39 specimen-to-electrode separation, or due to an actual, originating, atomic lo-
 40 cation on or near the specimen apex, are considered negligible. The detector
 41 surface is a flat, circular region roughly 50 mm in diameter. As a conse-

42 quence of geometry therefore, the flight length, L , and corresponding TOF
43 both depend upon where on the detector a detection event is registered; the
44 minimum flight length being 90 mm.

45 For analysis, the TOF data are transformed so that an m/z peak, for
46 example a 28-Da Si peak, will show the same apparent TOF over the entire
47 swing of V_{app} . The data are further transformed so that the (physically) flat
48 detector is transformed into a hypothetical hemispherical section so that an
49 ion’s apparent flight length will be independent of where it strikes the de-
50 tector. These transformations of APT data are generally called “voltage”
51 and “bowl” corrections [2, 3, 6]. Operationally, the voltage and bowl cor-
52 rections are performed on raw TOF data with IVAS software [9]. After the
53 voltage/bowl corrections, the TOF data are transformed to m/z data that
54 is then calibrated with IVAS software based on recognized locations of m/z
55 peaks for Si. Finally, the m/z data are transformed back to TOF data. We
56 then scale this TOF data so that its sample mean agrees with the sample
57 mean of the raw TOF data.

58 Due in part to thermal effects, the evaporation time of any ion and its
59 TOF are both random. Hence, ions with particular m/z values produce
60 temporal peaks that are broadened with respect to an idealized measurement
61 where an ion evaporates within a time interval of 8 ps after the arrival of
62 the laser pulse (which has a width of 8 ps). In the real situation, ions are
63 persistently emitted at times greater than 8 ps after the arrival of the pulse
64 since it takes time for the tip to cool below the threshold for field emission.
65 In this work, we focus on the problem of reconstruction of the m/z spectrum
66 of ions produced by laser energy deposits in a sample given a measured TOF
67 spectrum produced by the evaporated ions.

68 In [10], the authors review how practitioners often determine peak in-
69 tensities by ad hoc methods such as summing the number of counts in a
70 full-width-half region about the peak of the peak or fitting parametric mod-
71 els to peaks in a region of interest. However, as noted in [10], the assumed
72 parametric models are often inaccurate. In some situations, intensities of ion
73 species that produce overlapping peaks (typically in m/z -space) can be de-
74 termined with maximum likelihood methods given accurate estimates of their
75 local isotopic abundances [11, 12, 13]. However, this peak overlap approach
76 does not yield estimates of peaks on a fine scale over a broad range. Ideally,
77 one should reconstruct the intensity of any ion species given the probability
78 mass function (pmf) associated with the TOF peak of interest and the pmfs
79 associated with neighboring peaks that affect the TOF spectrum in the re-

80 gion of interest. See Appendix A for a definition of the pmf associated with
81 a particular peak.

82 In general, the TOF pmf for each ion of interest varies. For the case where
83 each TOF peak is described by a translated and scaled reference TOF pmf,
84 there are methods such as [10, 14, 15] to determine the TOF pmf. However,
85 as far we know, there are no published methods to determine TOF pmfs at
86 a fine scale over a broad range that vary from peak-to-peak.

87 Here, we determine TOF peak pmfs that that can vary from
88 peak-to-peak with a mixture model [16, 17] (a key tool in both Machine
89 Learning and Statistics) approach. We note that Vurpillot et al. [18] deter-
90 mined peak pmfs in m/z -space, but not TOF-space, with a mixture model
91 approach where each peak is assumed to have a Gaussian shape. However
92 the mixture model in [18] does not account for the contribution of a non-
93 constant background to m/z measurements. Further, as noted in [18], for
94 many atoms, peak shapes are non-Gaussian. In our approach, predicted peak
95 pmfs can have non-Gaussian shapes, and we account for the contribution of
96 a background to TOF measurements. We specify a collection of candidate
97 pmfs for each peak of interest and determine their associated intensities with
98 a maximum likelihood method. We determine candidate pmfs by a template
99 deformation method. Hence, our approach is conceptually similar to existing
100 deformable template methods [19, 20, 21, 22]. Based on intensities for each
101 peak of interest determined by a maximum likelihood method, we estimate
102 the pmf for the peak as the weighted average of its candidate pmfs. In a
103 proof-of-principle study, we apply our method to experimental data corre-
104 sponding to a silicon sample that was acquired at the National Institute of
105 Standards and Technology (NIST). In particular, we determine the intensi-
106 ties of three singly charged isotopes of silicon. We expect that our method
107 will apply to more complicated experimental cases as well. In Section 2, we
108 discuss experimental methods and our analysis methods. In Section 3, we
109 present our main results for a silicon sample. In Section 4, we discuss our
110 main results and some complimentary studies.

111 **2. Methods**

112 *2.1. Experimental method*

113 Atom probe tomography (APT) was performed with a LEAP 4000X Si
114 instrument (manufactured by CAMECA) operating in laser pulsing mode.
115 Voltage pulsing mode was not performed for this work. Data were collected

116 from a single Si microtip residing on a Sb-doped Si microtip coupon array,
117 which was also supplied by CAMECA. The conditions for data acquisition
118 were as follows.

119 (1) Pulsed laser: the laser wavelength, energy per pulse, and repetition
120 rate were 355 nm, 30 pJ, and 250 kHz, respectively. The pulse width is 8 ps.

121

122 (2) Analysis environment: the specimen stage was maintained at an in-
123 dicated temperature of 54 K under a vacuum of approximately 7.6×10^{-9} Pa.

124

125 (3) Detection rate of ions: the detection rate was 0.003 ions per laser
126 pulse.

127

128 (4) Span of specimen-electrode bias voltage: the data were collected over
129 a range of bias voltage spanning from 2.75 kV to 6.77 kV. The voltage is auto-
130 matically adjusted during data acquisition in order to maintain the specified
131 detection rate. Since the specimen tip is conical with a semi-hemispherical
132 cap, the apex diameter of the tip will increase as ions are field-evaporated
133 away. This increasing apex diameter necessitates continuously increasing the
134 bias voltage to maintain an electric field strength that provides the desired
135 evaporation rate of ions.

136

137 (5) Duration of data acquisition: data were accumulated over a run time
138 of 15.5 hours resulting in a data set composed of 40.5 million ions.

139

140 (6) Time-of-flight (TOF) mass spectrometer: the sample-to-detector flight
141 length was 90 mm.

142

143 (7) The specimen did not fracture during data acquisition. Since we were
144 only interested in TOF information and not 3D spatial reconstruction of the
145 data, we did not record the comparative apex diameters nor the heights of
146 the specimen tip before and after APT analysis.

147

148 *2.2. Mixture model and reconstruction*

149 Suppose that the mass-to-charge ratio for an ion is $r = m/z$ where m
150 and z are the mass and charge of the ion. Further, assume that the expected
151 number of ions with ratio r that produce counts in the observed TOF his-
152 togram is $\Theta(r)$. Assume that the conditional probability mass function for

153 the observed TOF histogram given r is $p_{theo}(t|r)$ where t is the midpoint of
 154 the bin. In general, $p_{theo}(t|r)$ varies with r . We model this theoretical pmf
 155 as a mixture of J candidate pmfs where the j th candidate pmf is $p(t|r, j)$, as
 156 follows:

$$p_{theo}(t|r) = \sum_j w_{r,j} p(t|r, j), \quad (2)$$

157 where the weights $w_{r,j}$ are non-negative and $\sum_j w_{r,j} = 1$.

158 Given our model assumptions, the expected number of counts in the k th
 159 TOF bin, $\mathbb{E}(n_k)$, is

$$\mathbb{E}(n_k) = \sum_r \sum_j w_{r,j} \Theta(r) p(t_k|r, j) + \lambda_b p_b(t_k), \quad (3)$$

160 where $p_b(t)$ is the pmf for the background process, and λ_b is the expected
 161 number of total background events. Rewriting Eq. 3, we get

$$\mathbb{E}(n_k) = \sum_r \sum_j \lambda(r, j) p(t_k|r, j) + \lambda_b p_b(t_k), \quad (4)$$

162 where

$$\lambda(r, j) = w_{r,j} \Theta(r). \quad (5)$$

163 We model the observed TOF histogram as a realization of a multivariate
 164 Poisson process where the associated Poisson random variables for the bins
 165 are independent. We determine estimates of the intensity parameters $\lambda(r, j)$
 166 and λ_b (which are Poisson parameters in the Poisson Likelihood model) with
 167 a maximum likelihood method based on the iterative Expectation Maximiza-
 168 tion (EM) algorithm [23]. This 1D-spectrum reconstruction method (not to
 169 be confused with the method for 3D-spatial reconstruction of APT data) is
 170 mathematically similar to reconstruction methods for emission tomography
 171 [24, 25, 26](see Appendix B). To avoid possible over-fitting of our mixture
 172 model to the observed TOF spectrum, we halt the EM algorithm with a cross-
 173 validation [27] method similar to those in [28, 29]. In our cross-validation
 174 analysis, we randomly split the observed TOF spectrum into training data
 175 and validation data by a binomial thinning method (see, for example, [30]).
 176 Based on the reconstruction determined from the training data, we compute

177 the conditional log-likelihood function of the validation data at each iteration
 178 of the EM algorithm. When this conditional likelihood first decreases, we halt
 179 the EM algorithm. Based on cross-validation analysis of multiple splits of
 180 the observed data, we determine the median iteration (i_{med}) at which the EM
 181 algorithm is halted. We then determine a reconstruction from the observed
 182 TOF spectrum after i_{med} iterations of the EM algorithm.

183 Given the EM algorithm estimate of each $\lambda(r, j)$, $\hat{\lambda}(r, j)$, we estimate $\Theta(r)$
 184 and $w_{r,j}$ as $\hat{\Theta}(r)$ and $\hat{w}_{r,j}$ where

$$\hat{\Theta}(r) = \sum_j \hat{\lambda}(r, j), \quad (6)$$

185 and

$$\hat{w}_{r,j} = \frac{\hat{\lambda}(r, j)}{\hat{\Theta}(r)}. \quad (7)$$

186 2.3. Construction of candidate pmfs

187 Based on an empirical estimate of the TOF pmf for the $^{28}\text{Si}^{1+}$ peak, we
 188 construct a reference pmf, $p_{ref}(t)$, where the first bin corresponds to the start
 189 time of the peak. For each of the other two peaks, we translate the reference
 190 pmf to nine possible locations in the vicinity of the nominal start time of the
 191 peak of interest. For each translation, we generate seven candidate pmfs by
 192 distorting the reference pmf as follows:

$$p_{distort}(t, \beta_i) = C_i p_{ref}^{\beta_i}(t), \quad (8)$$

193 where $\beta_i = 0.4 + (i - 1) * 0.2$ for $i = 1, 2, \dots, 7$, and C_i is a normalization
 194 constant that ensures the pmf sums to one. For the $^{29}\text{Si}^{1+}$ and $^{30}\text{Si}^{1+}$ peaks,
 195 we determine 63 intensity parameters for each. We determine an additional
 196 intensity parameter for the $^{28}\text{Si}^{1+}$ peak, and another for the background
 197 process. In all, we determine 128 intensity parameters.

198 3. Results

199 Based on the fully corrected TOF data (see Introduction for details about
 200 the corrections), we compute a TOF histogram where the width of each bin
 201 is 0.1 ns. We analyze a subset corresponding to TOF values in the interval
 202 [417.95, 439.95] ns (see Figure 1) corresponding to singly-charged isotopes of
 203 silicon.

204 We estimate a reference pmf based on analysis of the TOF spectrum as-
 205 sociated with the $^{28}\text{Si}^{1+}$ peak. We determine a flat background by averaging
 206 TOF values in the interval [436,440] ns. The estimated background (98.7) is
 207 subtracted from the observed TOF spectrum. We fit an exponential model
 208 to the tail of the background-corrected $^{28}\text{Si}^{1+}$ peak (see Figure 2a). The ob-
 209 served values in the region of the exponential fit (thick red line in Figure 2a)
 210 are replaced by the exponential model prediction. We normalize the extrap-
 211 olated version of the Figure 2a spectrum to get a reference pmf that sums to
 212 1 for relative TOF values up to 1000 ns. We generate candidate pmfs from
 213 this reference pmf according to Eq. 8. We normalize each candidate pmf
 214 so that it sums to 1 (see Figure 2b). We determine a nominal start time
 215 for the $^{29}\text{Si}^{1+}$ peak as the nominal start time of the $^{28}\text{Si}^{1+}$ peak plus the
 216 difference between the TOF time where the $^{29}\text{Si}^{1+}$ takes its maximum value
 217 and TOF time where the $^{28}\text{Si}^{1+}$ peak takes its maximum value. A similar
 218 method yields a nominal start time for the $^{30}\text{Si}^{1+}$ peak.

219 Given n counts in a particular bin of the TOF histogram, the number of
 220 thinned counts in the same particular bin of the training data, n_{train} , is a
 221 realization of Binomial random variable with parameters n and p . Here, we
 222 set $p = 0.8$. Given n_{train} , the number of thinned counts in the same particular
 223 bin in the validation data is $n - n_{train}$. For each of five sets of training and
 224 validation data, we determine estimates of the 128 Poisson parameters in
 225 our model (see Section 2.2) from the training data with the EM algorithm.
 226 At each iteration of the EM algorithm, we compute the conditional log-
 227 likelihood of the validation data based on the Poisson parameters determined
 228 from the training data. When this conditional log-likelihood first decreases,
 229 we halt the EM algorithm. Based on analysis of the five sets of training
 230 and validation data, we determine the median iteration at which the EM
 231 algorithm is halted to be $i_{med} = 33\ 753$. From the original TOF histogram,
 232 we determine Poisson parameters with EM algorithm after i_{med} iterations.
 233 In our implementation of the EM algorithm, the initial value of the Poisson
 234 parameter for each candidate pmf (excluding the background pmf) is the
 235 sum of counts in the observed histogram divided by 127. The initial value
 236 of the background intensity is $99 * n_{bin} * f$ where f is 0.8 for training data
 237 produced by thinning, f is 1 for the observed TOF histogram, and n_{bin} is
 238 the number of bins (221) in each TOF histogram.

239 In Figures 3 and 4, for the $^{29}\text{Si}^{1+}$ and $^{30}\text{Si}^{1+}$ peaks, for each candidate
 240 translation, we display the estimated intensity for each of seven candidate
 241 pmfs. In Figure 5a, for the $^{29}\text{Si}^{1+}$ peak, for each candidate translation, we

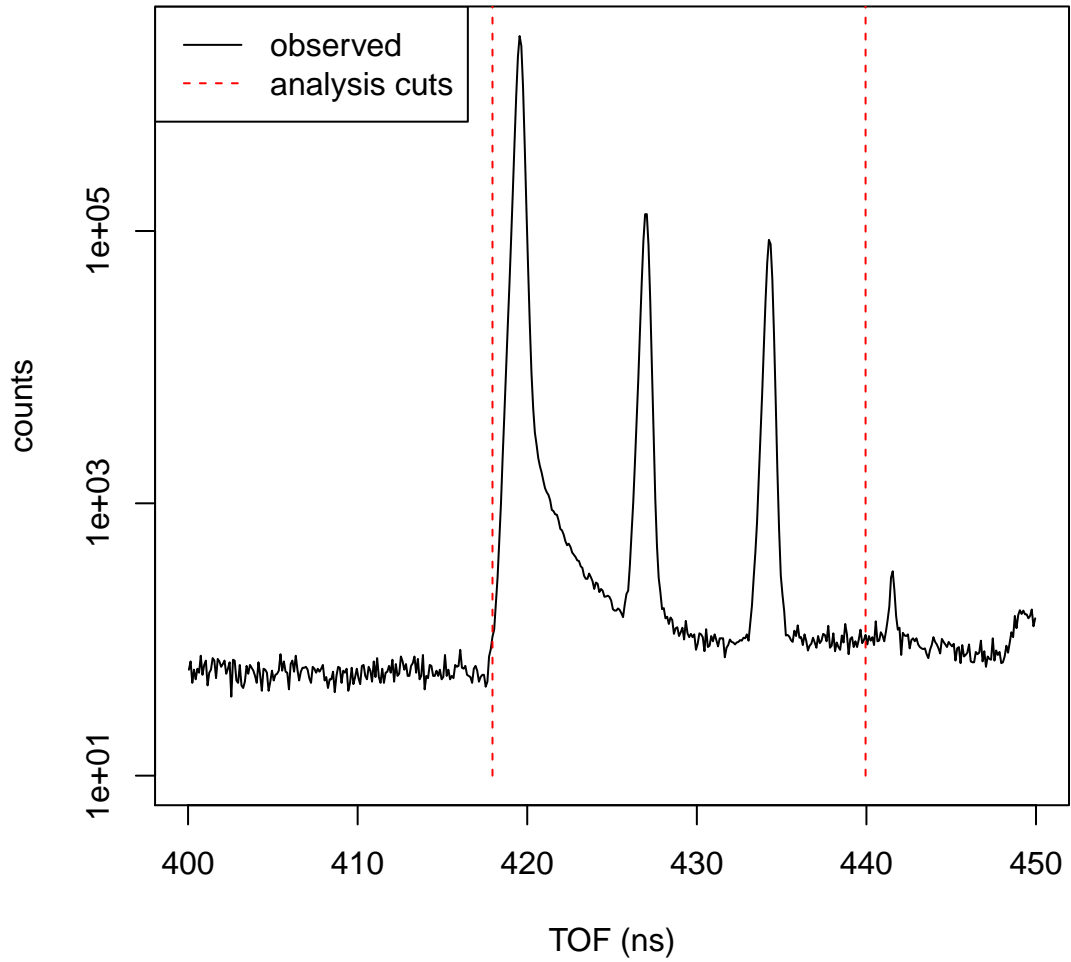


Figure 1: Observed TOF spectra for a silicon sample. The three peaks inside the analysis region (from 419.95 ns to 439.95 ns) are due to singly charged isotopes of silicon— $^{28}\text{Si}^{1+}$, $^{29}\text{Si}^{1+}$, and $^{30}\text{Si}^{1+}$. There are 221 bins with width of 0.1 ns. From count data corresponding to TOF values in the interval [436,440] ns, we determine the background to be 98.7 counts.

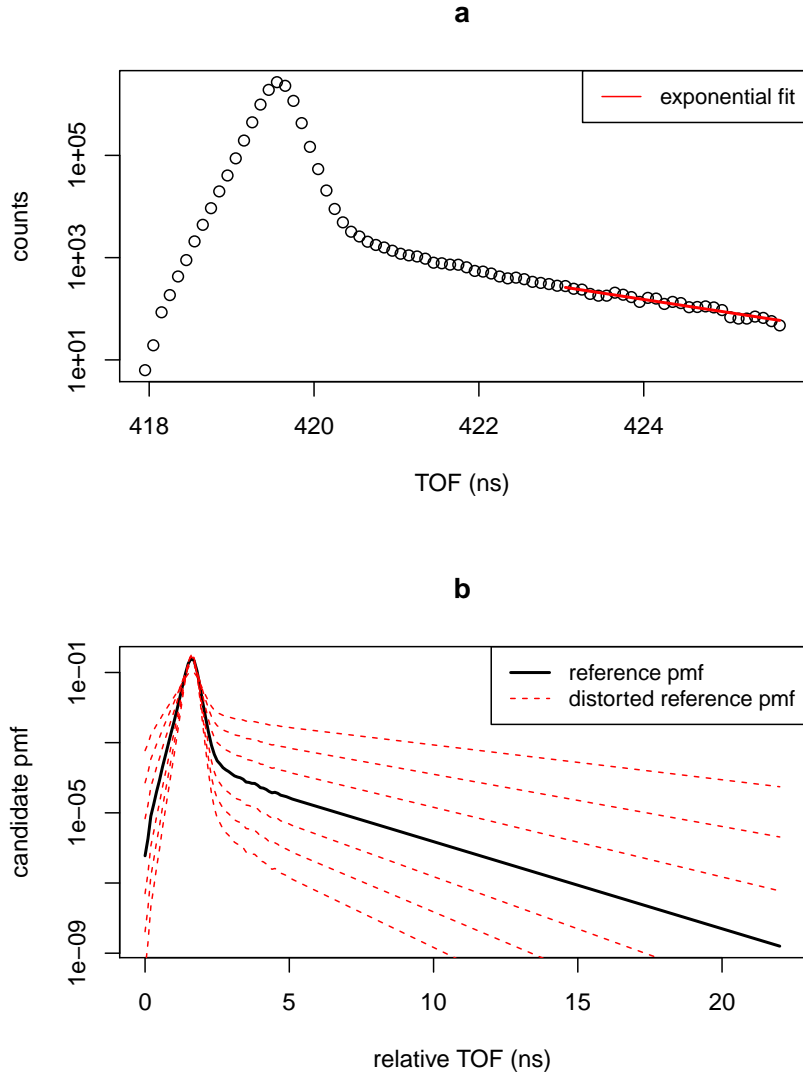


Figure 2: In (a), we fit an exponential model to the background-corrected data produced by the $^{28}\text{Si}^{1+}$ peak. (The fitted value of the mean lifetime of the exponential and its standard uncertainty are $\tau = 1.75$ ns and 0.09 ns.) Based on the estimated exponential model parameters, we extrapolate to arbitrarily large values of TOF. Based on the extrapolated curve in (a), we determine a reference pmf. In (b), based on Eq. 8, we generate candidate pmfs by deforming the reference pmf.

242 plot the sum of the seven intensities. In Figure 5b, for the $^{30}\text{Si}^{1+}$ peak, for
 243 each candidate translation, we plot the sum of the seven intensities. Based on
 244 Eq. 6, the overall intensity for the $^{29}\text{Si}^{1+}$ peak is the sum of the 63 intensities
 245 shown in Figure 3. Similarly, the overall intensity for the $^{30}\text{Si}^{1+}$ peak is the
 246 sum of the 63 intensities shown in Figure 4.

247 In Figure 6, we compare observed and predicted TOF spectra and show
 248 standardized residuals. In Figure 7, we compare the contribution of each
 249 peak to the overall TOF spectrum. In Figure 8, we compare the shapes of
 250 the pmfs for the three peaks. In Table 1, we list estimated intensities for each
 251 peak and the estimated isotopic abundance associated with each isotope of
 252 silicon.

Table 1: Below, we show estimates of intensities and associated isotopic abundances for
 for $^{28}\text{Si}^{1+}$, $^{29}\text{Si}^{1+}$ and $^{30}\text{Si}^{1+}$ determined at 33 753 iterations of the iterative EM algorithm.
 This halting iteration is determined by cross-validation. The theoretical values for isotopic
 abundances of ^{28}Si , ^{29}Si , and ^{30}Si are 0.9222, 0.0469 and 0.0309. For iterations of EM
 algorithm greater than 10 000 iterations, results agree to the number of significant digits
 shown below.

peak	intensity	isotopic abundance
$^{28}\text{Si}^{1+}$	1.05549E+07	0.92121
$^{29}\text{Si}^{1+}$	5.517E+05	0.04815
$^{30}\text{Si}^{1+}$	3.511E+05	0.03064

253 4. Discussion

254 4.1. Reconstruction method and Poisson model

255 As discussed earlier, we assume that the observed TOF histogram is a
 256 realization of a multivariate Poisson process. Visual inspection of the 146
 257 standardized residuals computed according to the Poisson model (see Fig-
 258 ure 6b) at times beyond the nominal start time of the $^{29}\text{Si}^{1+}$ peak suggests
 259 that our Poisson model and reconstruction method are reasonable. Further,
 260 the fact that the root-mean-square (RMS) value of these 146 standardized
 261 residuals is 1.08 suggests that our Poisson model and reconstruction method
 262 are reasonable. We note that standardized residuals for early times of the
 263 $^{28}\text{Si}^{1+}$ peak are generally smaller in magnitude compared to the magnitude
 264 of standardized residuals at later times. We attribute this to the fact that
 265 the assumed pmf values for the $^{28}\text{Si}^{1+}$ peak at early times are proportional
 266 to background-corrected observed values.

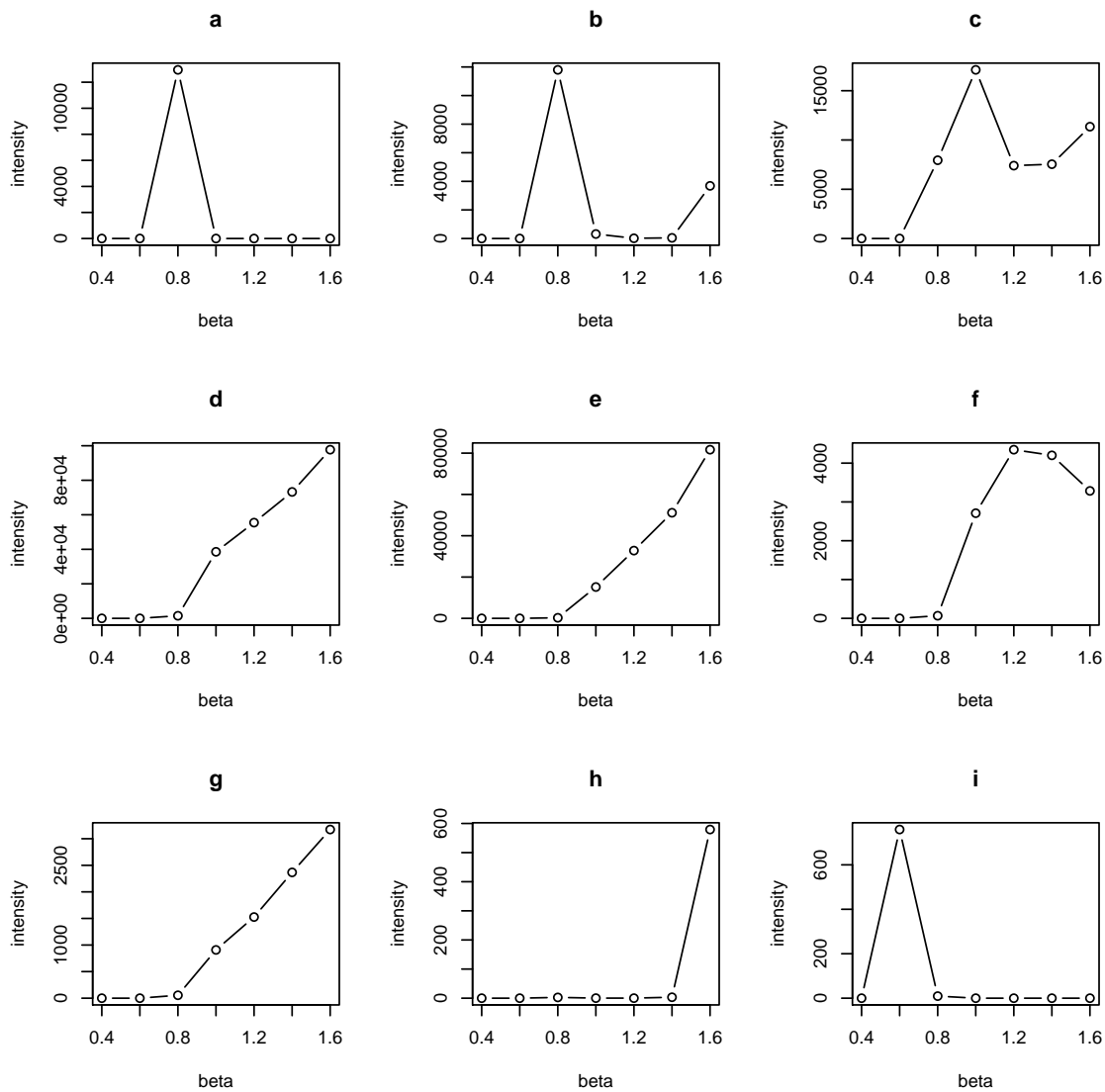


Figure 3: For the $^{29}\text{Si}^{1+}$ peak, we show intensities of candidate pmfs determined with the EM algorithm after 33 753 iterations (determined by cross-validation). For each candidate translation about the nominal start time of the $^{29}\text{Si}^{1+}$ peak, (-0.4, 0.3, 0.2, 0.1, 0, 0.1, 0.2, 0.3, 0.4) ns, we show intensities for the seven candidate pmfs in plots (“a”, “b”, “c”, “d”, “e”, “f”, “g”, “h”, “i”).

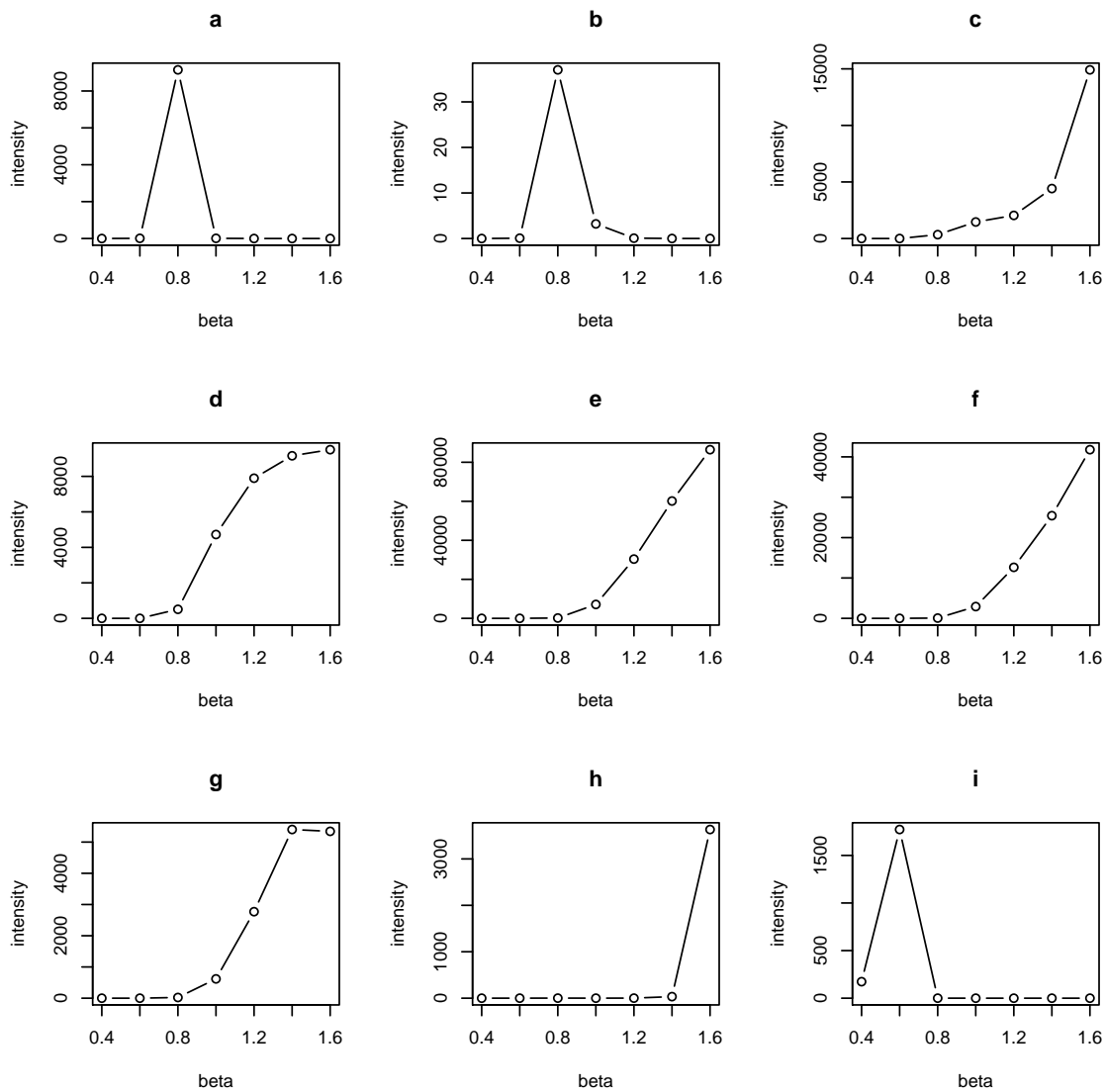


Figure 4: For the $^{30}\text{Si}^{1+}$ peak, we show intensities of candidate pmfs determined with the EM algorithm after 33 753 iterations (determined by cross-validation). For each candidate translation about the nominal start time of the $^{30}\text{Si}^{1+}$ peak, (-0.4, 0.3, 0.2, 0.1, 0, 0.1, 0.2, 0.3, 0.4) ns, we show intensities for the seven candidate pmfs in plots (“a”, “b”, “c”, “d”, “e”, “f”, “g”, “h”, “i”).

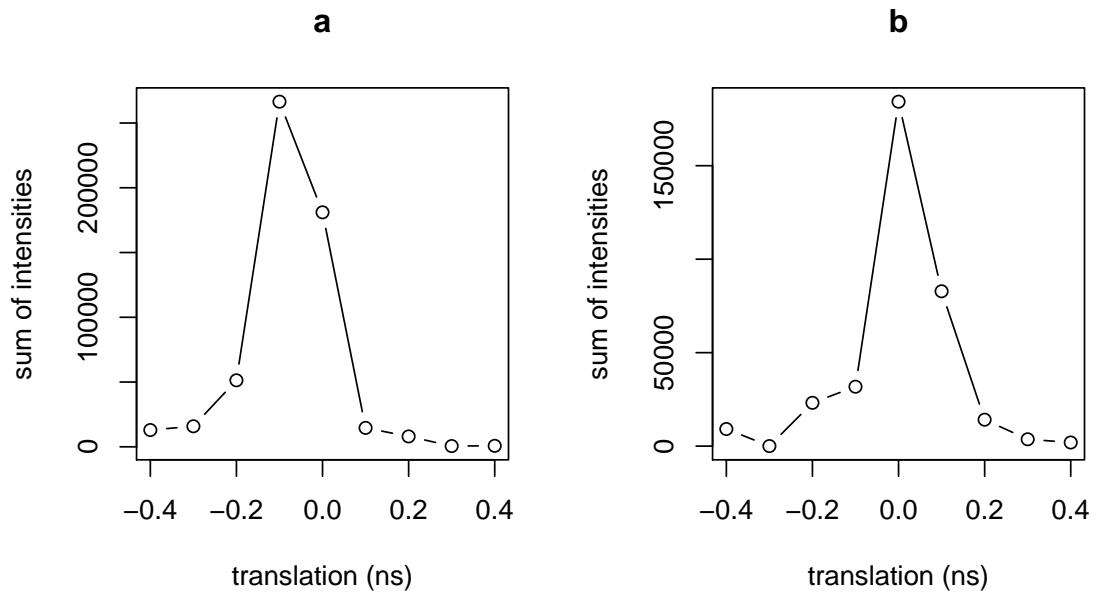


Figure 5: In (a) and (b), we show the sum of the intensities of the seven candidate pmfs for each candidate translation for the $^{29}\text{Si}^{1+}$ peak and the $^{30}\text{Si}^{1+}$ peak respectively.

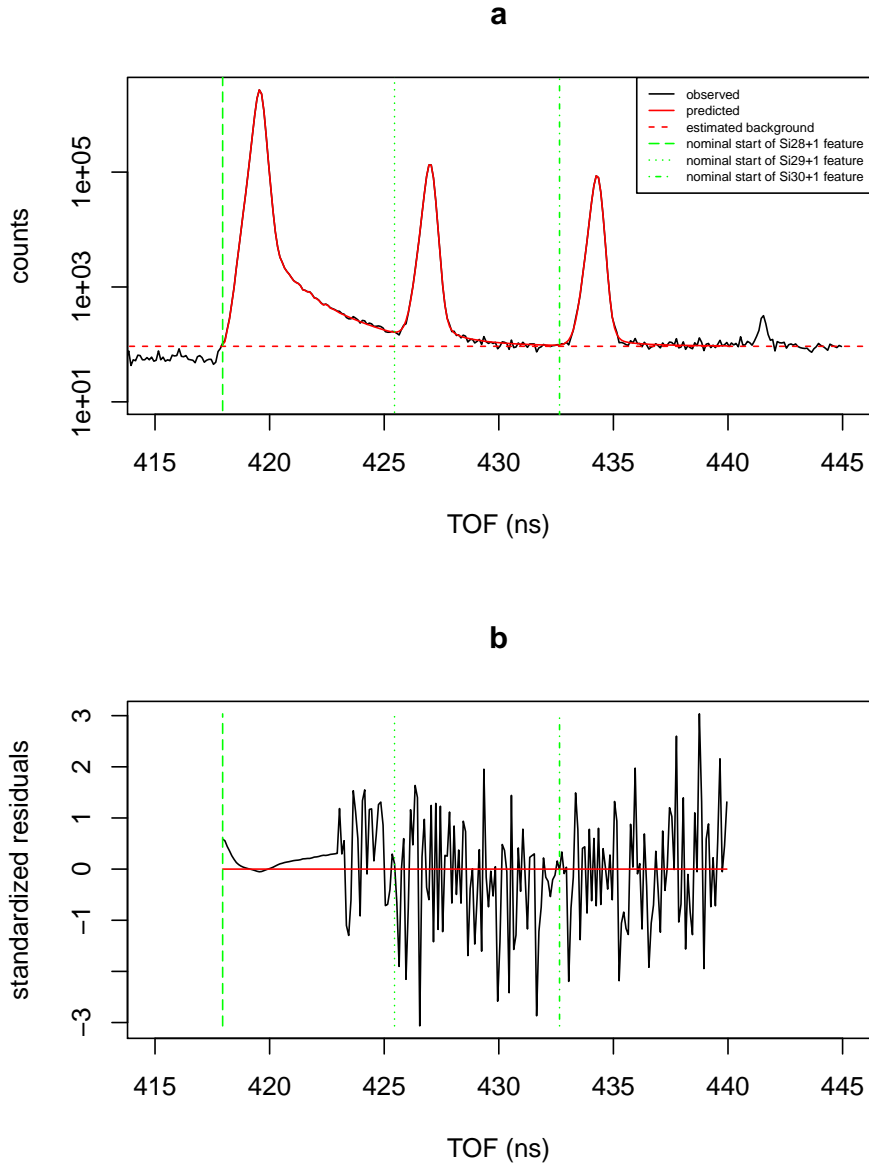


Figure 6: (a): Observed and predicted TOF spectra. The EM algorithm determines the background to be 92.7 counts for results in the analysis region (see Figure 1). (b): Standardized residuals according to our Poisson model (difference between observed and predicted values divided by the estimated standard deviation of the observation (the square root of the predicted value)). The root-mean-square (RMS) value of the 146 standardized residuals corresponding to TOF values as large or larger than the nominal start time of the $^{29}\text{Si}^{1+}$ peak is 1.08.

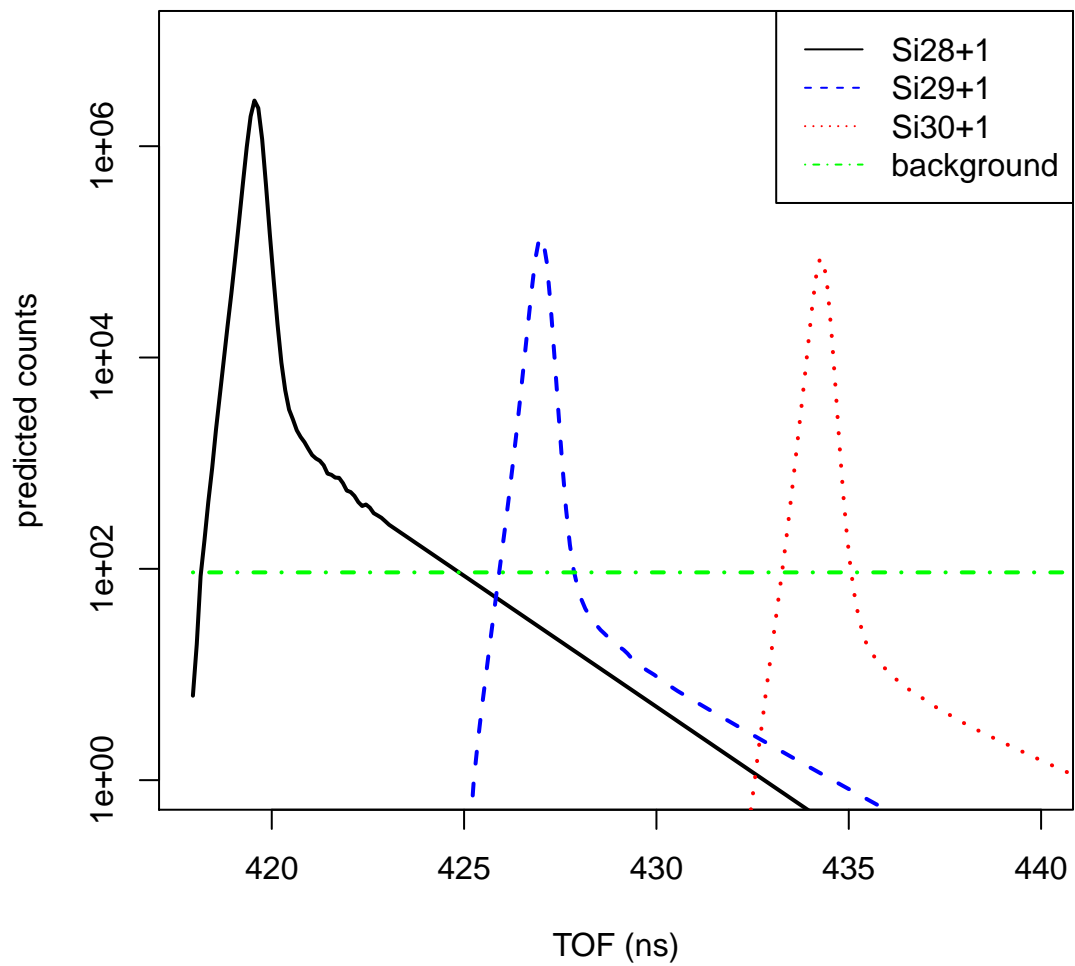


Figure 7: Contributions of the ($^{28}\text{Si}^{1+}$, $^{29}\text{Si}^{1+}$, $^{30}\text{Si}^{1+}$) peaks and the background (92.7 counts) to the overall TOF spectrum predicted by our mixture model analysis.

267 Our reconstruction method conserves counts. That is, the sum of ob-
 268 served counts in all bins and the sum of predicted counts in all bins agree.

269 In our model, the observed number of counts in the k th TOF histogram
 270 bin, n_k , is a realization of a Poisson process with expected value λ_k . Given
 271 that the sum of observed counts in all K bins is n , it is well-known (see, for
 272 example, [40]) that (X_1, X_2, \dots, X_K) is a multinomial random variable with
 273 the following distribution (probability mass function)

$$\Pr(X_1 = n_1, X_2 = n_2, \dots, X_K = n_K | X = n) = \frac{n!}{n_1! n_2! \dots n_K!} \prod_{k=1}^K p_k, \quad (9)$$

274 where $X = X_1 + X_2 + \dots + X_K$, and

$$p_k = \frac{\lambda_k}{\lambda_1 + \lambda_2 + \dots + \lambda_K}. \quad (10)$$

275 4.2. Stability of estimates

276 In a stability study, we determine results with the EM algorithm for
 277 iterations ranging from 1000 to 4 million. As discussed earlier, the overall
 278 intensity estimates for the $^{29}\text{Si}^{1+}$ peak and the $^{30}\text{Si}^{1+}$ peak are the sum of
 279 estimated intensities determined for each of 63 candidate pmfs (see Eq. 6).
 280 After 10 000 iterations, the estimate of the intensity of the $^{29}\text{Si}^{1+}$ peak and
 281 the estimate of the $^{30}\text{Si}^{1+}$ peak are both stable to within four significant
 282 digits. After 10 000 iterations, the estimate of the intensity of the $^{28}\text{Si}^{1+}$
 283 peak is stable to within six significant digits. For either the $^{29}\text{Si}^{1+}$ or $^{30}\text{Si}^{1+}$
 284 peak, in general, the intensity associated with any one of the 63 candidate
 285 pmfs is not stable to within four digits after 10 000 iterations. For instance,
 286 for the $^{29}\text{Si}^{1+}$ peak, the ratios of the intensity estimates for candidate pmfs at
 287 4 million iterations and i_{med} iterations range from 2.7E-320 to 14.3. For the
 288 $^{30}\text{Si}^{1+}$ peak, the ratios range from 1.5E-321 to 6315. One expects the average
 289 of the 63 intensities for each peak to be better determined than any one of
 290 the 63 estimates corresponding to that peak. In medical imaging studies, a
 291 similar phenomenon occurs when EM algorithm estimates of a large number
 292 of spatially varying Poisson parameter estimates are averaged within regions
 293 of interest (for example, see [31, 32]). Estimates of the predicted $^{29}\text{Si}^{1+}$
 294 and $^{30}\text{Si}^{1+}$ peaks. at any particular TOF value are less stable than overall
 295 intensities of these peaks (see Figure 9).

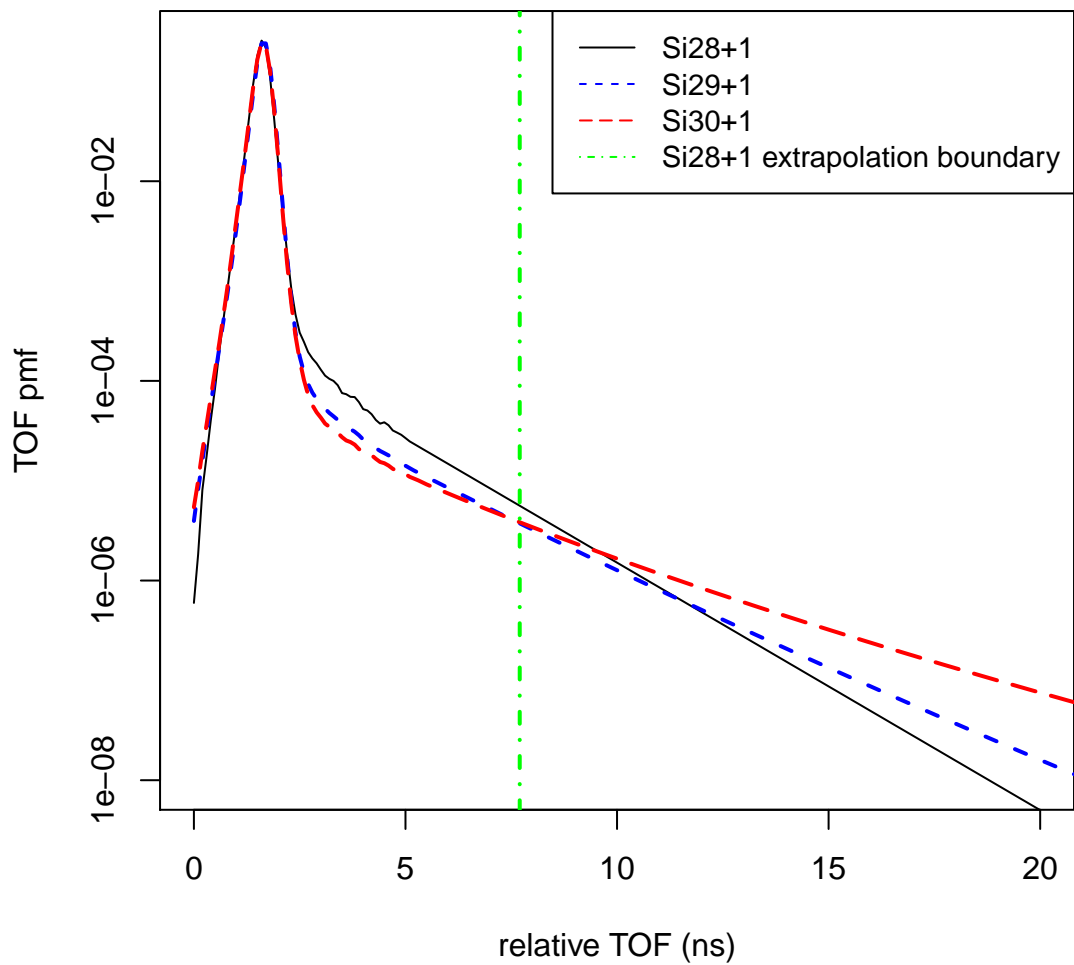


Figure 8: Comparison of aligned estimated probability mass functions for the ($^{28}\text{Si}^{1+}$, $^{29}\text{Si}^{1+}$, $^{30}\text{Si}^{1+}$) peaks determined with EM algorithm after 33 753 iterations (determined by cross-validation).

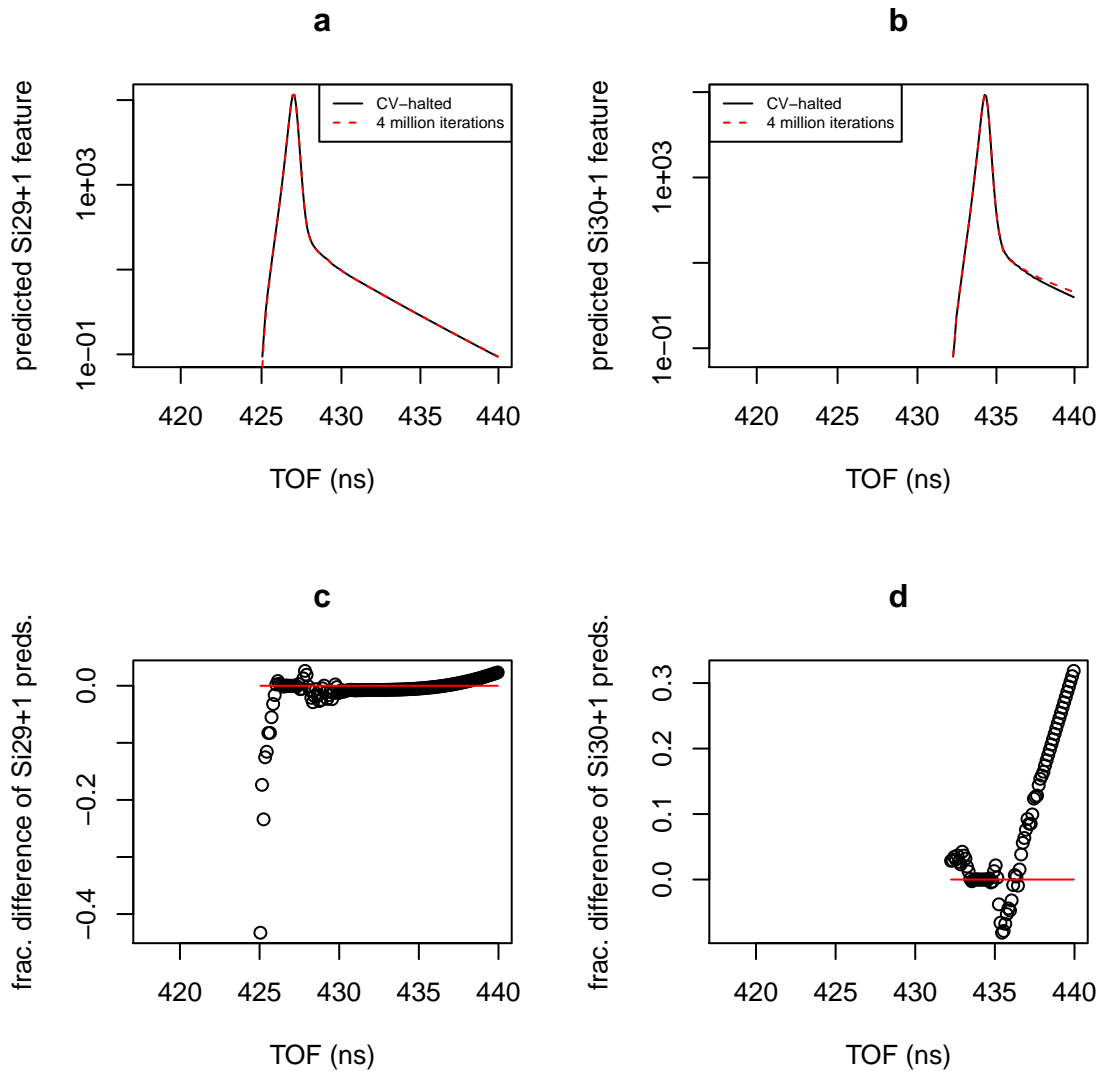


Figure 9: (a) Comparison of predicted $^{29}\text{Si}^{1+}$ peak determined by the EM algorithm after 33 753 iterations (determined by cross-validation) and 4 million iterations. (b) Comparison of predicted $^{30}\text{Si}^{1+}$ peak determined by the EM algorithm after 33 753 iterations and 4 million iterations. (c) Fractional difference between predictions shown in (a). (d) Fractional difference between predictions shown in (b).

296 *4.3. Initial values of Poisson parameter estimates*

297 When the initial values of the Poisson parameter estimates in the EM
298 algorithm method are equated to the square root of the values assumed in
299 our main study, for i_{med} iterations, the estimated value of the estimated 128
300 Poisson parameters are similar. The ratios of the new and original estimates
301 of the Poisson parameters vary from 0.93 to 1.01, and the median ratio is
302 1.00. The fractional differences between the predicted value of the TOF
303 spectrum with new initial value relative to the predicted value determine
304 with the original initial values ranges from $-2.1E-05$ to $1.3E-05$ (over all bins).
305 Further, the estimates of the intensities (determined with the original and
306 revised choice of initial values of the Poisson parameters) for the $^{28}\text{Si}^{1+}$ $^{29}\text{Si}^{1+}$
307 and $^{30}\text{Si}^{1+}$ peaks agreed to within nine significant digits, six significant digits,
308 and six significant digits respectively.

309 *4.4. Background jumps*

310 Our comparison of predicted and observed TOF spectra shown in Figure
311 6a strongly suggests that the background rate before the start of the $^{28}\text{Si}^{1+}$
312 peak is less than the background rate after the start of the $^{28}\text{Si}^{1+}$ peak. The
313 source of this jump could be an ion feedback (IFB) artifact produced by
314 micro-channel plates (MCPs) in the single ion detection systems [34, 35, 36,
315 37, 38, 10].

316 *4.5. Peak-to-peak pmf variation*

317 Our results appear consistent with the hypothesis that the TOF pmfs vary
318 among the three peaks (see Figure 8). As a caveat, the source of this apparent
319 variation among the pdfs could be due to the physics of ion evaporation or
320 systematic effects associated with multiple-hit events [33], IFB artifacts, or
321 perhaps other instrumental effects.

322 *4.6. Extrapolation model*

323 In our analysis, we model the upper tail of the $^{28}\text{Si}^{1+}$ peak as an expo-
324 nential (see Figure 2). While this model appears to be reasonable, we can
325 not rule out the possibility that a non-exponential model is more accurate.

326 *4.7. Application to other experimental data*

327 In our study, the peaks due to $^{28}\text{Si}^{1+}$, $^{29}\text{Si}^{1+}$, $^{30}\text{Si}^{1+}$, overlap, but peaks
328 that are not in analysis region appear to not significantly overlap with the
329 silicon peaks. If other peaks beyond the analysis had significantly overlapped
330 with the silicon peaks, we should account for the contribution of these peaks.
331 We believe that our mixture model approach is flexible enough to model the
332 contributions of such peaks. For other applications, the method to generate
333 candidate pmfs may be different than the method presented here. For in-
334 stance, for peaks with shapes that vary more dramatically (see, for example,
335 [41]) than the peaks considered in this work, the number of candidate pmfs
336 for each peak might be larger than assumed in this work and analysis could
337 be more computer intensive. Due to the flexibility of our method, it may
338 be possible to determine candidate peak pmfs from deformations of analyti-
339 cal models. Scaling up the mixture model analysis to a larger subset of the
340 observed TOF spectrum is a worthwhile topic for future research.

341 *4.8. Uncertainties*

342 Quantifying uncertainties for results is a worthy topic for future research.

343 **5. Conclusions**

344 We presented a mixture model method to determine TOF pmfs that can
345 vary from peak-to-peak. In a proof-of-principle study based on experimental
346 data, we demonstrated the feasibility of our method by determining a TOF
347 pmf, intensity and isotopic abundance for each of three singly charged iso-
348 topes of silicon. How well our method works for more complicated examples
349 is a topic for future study.

350
351 **Appendix A: Probability mass function**

352
353 Consider an ideal experiment where time-of-flight measurements are pro-
354 duced by ions of the same species. Further, assume that each time-of-flight
355 measurement is a realization of a random variable T . Suppose that we rep-
356 resent the TOF spectrum associated with this ion species as a histogram
357 with K bins where K can be arbitrarily large. Further, suppose that the
358 probability that any realization of T falls in the k th bin is $f(k)$. Since T is a
359 random variable, the bin index is a random variable $g(T)$ where g is a known

360 function that determines the bin index given the realization of T . We denote
 361 the probability that $g(T) = k$ as

$$\Pr(g(T) = k) = f(k), \quad (\text{A.1})$$

362 where $k = 1, 2, \dots, K$. Given that $\sum_{k=1}^K f(k) = 1$, $f(k)$ is a probability mass
 363 function. In this work, $K = 10\,000$ for each pmf of interest. For this large
 364 value of K , we assume that $\sum_{k=1}^K f(k) = 1$. As discussed in many references
 365 (see, for example, [39]), the probability mass function is sometimes referred
 366 to as a discrete density function or probability function. In our study, we
 367 seek to determine pmfs associated with different ion species that can vary
 368 from species-to-species, as well as the intensity for each peak.

369 **Appendix B: EM algorithm applied to APT**

370
 371
 372 Following the nomenclature in [24] and [25], we assign an integer index b
 373 to each of the Poisson parameters in our model (see Section 2.2) (where $b =$
 374 $1, 2, \dots, B$) and an integer index d (where $d = 1, 2, \dots, D$) to each bin of the TOF
 375 histogram. In our study, $B = 128$ and $D = 221$. The Poisson likelihood of
 376 the observed TOF histogram $\mathbf{n} = (n_1, n_2, \dots, n_D)$ given the associated Poisson
 377 parameters (intensities) for the candidate pmfs $\lambda = (\lambda(1), \lambda(2), \dots, \lambda(B))$ is

$$L(\mathbf{n}|\lambda) = \prod_{d=1}^D \frac{e^{-\lambda^*(d)} \lambda^*(d)^{n(d)}}{n(d)!}, \quad (\text{B.1})$$

378 where the mean of $n^*(d)$ is

$$\lambda^*(d) = \sum_{b=1}^B \lambda(b) p(b, d), \quad (\text{B.2})$$

379 and $p(b, d)$ is the probability that the observed TOF will fall into the d th
 380 TOF histogram bin predicted by the b th pmf.

381
 382 The EM algorithm is an iterative scheme for determination of maximum
 383 likelihood estimates of the Poisson parameters. The b th Poisson parameter,
 384 $\hat{\lambda}$, is updated as follows:

$$\hat{\lambda}(b)^{new} = \frac{\hat{\lambda}(b)^{old}}{p(b, \cdot)} \sum_{d=1}^D \frac{n(d) p(b, d)}{\sum_{b'=1}^B \hat{\lambda}(b')^{old} p(b', d)}, \quad (\text{B.3})$$

385 where $p(b, \cdot) = \sum_{d=1}^D p(b, d)$. For more technical details on how to derive Eq.
386 B.3, see Section 2 of [25]. Shepp and Vardi [24] proved that, after each itera-
387 tion of the EM algorithm, the likelihood of the data never decreases provided
388 that the initial estimates are positive.

389

390 In our study, we estimate the expected background contribution to each
391 TOF measurement. For the case where the expected background contribu-
392 tion is treated as known, one can derive an EM algorithm (see, for example,
393 Fessler and Hero [42]) similar to Eq. B.3

394

395 **Acknowledgments** We thank B. Caplins, F. Meisenkothen and A. Pin-
396 tar of NIST for helpful comments. Contributions of staff of NIST, an agency
397 of the US Government are not subject to copyright in the US. Certain com-
398 mercial equipment, instruments, or materials are identified in this paper to
399 foster understanding. Such identification does not imply recommendation
400 or endorsement by the National Institute of Standards and Technology, nor
401 does it imply that the materials or equipment identified are necessarily the
402 best available for the purpose.

403

404 **Author Contributions** Norman Sanford acquired the raw APT mea-
405 surements at NIST and processed them to yield the TOF measurements an-
406 alyzed here. He also contributed text on the APT method and experimental
407 details. Kevin Coakley developed the reconstruction method and associated
408 software to reconstruct TOF measurements. He also analyzed results and
409 was the primary author of the manuscript. Both authors have approved the
410 final article for publication.

411

412 **Declaration of Competing Interest** The authors declare that they
413 have no known competing financial interests or personal relationships that
414 could have appeared to influence the work reported in this paper.

415

416 **Funding** This research did not receive any specific grant from funding
417 agencies in the public, commercial, or not-for-profit sectors.

418

419 **Data Availability** Data underlying the results presented in this paper
420 are not publicly available at this time but may be obtained from the authors
421 upon reasonable request.

422

References

- 423
- 424 [1] B. Gault, M. P. Moody, J. M. Cairney, S. P. Ringer, Atom Probe Mi-
425 croscopy, Vol. 160 of Springer Series in Materials Science, Springer New
426 York, New York, NY, 2012.
- 427 [2] D. J. Larson, T. J. Prosa, R. M. Ulfing, B. P. Geiser, T. F. Kelly, Local
428 Electrode Atom Probe Tomography, Springer New York, New York, NY,
429 2013.
- 430 [3] M. K. Miller, R. G. Forbes, Atom-Probe Tomography, Springer US,
431 Boston, MA, 2014.
- 432 [4] C. Pareige, W. Lefebvre-Ulrikson, F. Vurpillot, X. Sauvage, Atom Probe
433 Tomography, Elsevier, Cambridge, MA, 2016.
- 434 [5] A. Devaraj, D.E. Perea, J. Liu, L.M. Gordon, T.J. Prosa, P. Parikh,
435 D.R. Diercks, S. Meher, R.P. Kolli, Y.S. Meng and S. Thevuthasan
436 (2018). Three-dimensional nanoscale characterisation of materials by
437 atom probe tomography. International Materials Reviews, 63(2), pp.68-
438 101.
- 439 [6] B.W. Caplins, P.T. Blanchard, A.N. Chiamonti, D.R. Diercks, L.
440 Miaja-Avila, and N.A. Sanford (2020). An algorithm for correcting sys-
441 tematic energy deficits in the atom probe mass spectra of insulating
442 samples. Ultramicroscopy, 213, 112995.
- 443 [7] N.A. Sanford, P.T. Blanchard, M. Brubaker, K.A. Bertness, A. Roshko,
444 J.B. Schlager, R. Kirchhofer, D. R. Diercks, B. Gorman, (2014). Laser-
445 assisted atom probe tomography of MBE grown GaN nanowire het-
446 erostructures. *physica status solidi (c)*, 11(34), 608-612. [https://doi.
447 org/10.1002/pssc.201300579](https://doi.org/10.1002/pssc.201300579)
- 448 [8] Q. Zhang, B. Klein, N.A. Sanford, et al. Comparative Apex Electrostat-
449 ics of Atom Probe Tomography Specimens. *Journal of Elec Materi* 50,
450 30223029 (2021). <https://doi.org/10.1007/s11664-021-08932-6>
- 451 [9] <https://www.cameca.com/service/software/ivas>
- 452 [10] F. Meisenkothen, D.V. Samarov, I. Kalish, and E.B. Steel (2020). Ex-
453 ploring the accuracy of isotopic analyses in atom probe mass spectrom-
454 etry. *Ultramicroscopy*, 216, p.113018.

- 455 [11] A.J. London, D. Haley and M.P. Moody (2017). Single-ion deconvolu-
456 tion of mass peak overlaps for atom probe microscopy. *Microscopy and*
457 *Microanalysis*, 23(2), pp.300-306.
- 458 [12] A.J. London (2019). Quantifying uncertainty from mass-peak overlaps
459 in atom probe microscopy. *Microscopy and Microanalysis*, 25(2), pp.378-
460 388.
- 461 [13] J. Keutgen, A.J. London and O. Cojocaru-Mirdin (2021). Solving peak
462 overlaps for proximity histogram analysis of complex interfaces for atom
463 probe tomography data. *Microscopy and Microanalysis*, 27(1), pp.28-35.
- 464 [14] L.J.S. Johnson, M. Thuvander, K. Stiller, M. Oden, L. Hultman (2013),
465 Blind deconvolution of time-of-flight mass spectra from atom probe Ul-
466 tramicroscopy, 132, pp. 60-64.
- 467 [15] F. Meisenkothen, D. Samarov, M. McLean, I. Kalish, and E. Steel
468 (2021). Adaptive Peak Fitting for Isotope Analysis via Atom Probe Mass
469 Spectrometry. *Microscopy and Microanalysis*, 27(S1), pp.176-177.
- 470 [16] G.J. McLachlan, S.X. Lee and S.I. Rathnayake (2019). Finite mixture
471 models. *Annual review of statistics and its application*, 6, 355-378.
- 472 [17] G.J. McLachlan and D. Peel (2004). *Finite Mixture Models*, Wiley Series
473 in Probability and Statistics.
- 474 [18] F. Vurpillot, C. Hatzoglou, B. Radiguet, G. Da Costa, F. Delaroche and
475 F. Danoix (2019). Enhancing element identification by expectationmax-
476 imization method in atom probe tomography. *Microscopy and Micro-*
477 *analysis*, 25(2), pp.367-377.
- 478 [19] Y. Amit, U. Grenander, and M. Piccioni (1991). Structural image
479 restoration through deformable templates. *Journal of the American Sta-*
480 *tistical Association*, 86(414), 376-387.
- 481 [20] A.L. Yuille, P.W. Hallinan and D.S. Cohen (1992). Feature extraction
482 from faces using deformable templates. *International journal of com-*
483 *puter vision*, 8(2), 99-111.
- 484 [21] A.K. Jain, Y. Zhong and S. Lakshmanan (1996). Object matching us-
485 ing deformable templates. *IEEE Transactions on pattern analysis and*
486 *machine intelligence*, 18(3), pp.267-278.

- 487 [22] A.V. Dalca, M. Rakic, J. Guttag, and M.R. Sabuncu (2019). Learning
488 conditional deformable templates with convolutional networks. arXiv
489 preprint arXiv:1908.02738.
- 490 [23] A.P. Dempster, N.M. Laird and D.B. Rubin (1977). Maximum likeli-
491 hood from incomplete data via the EM algorithm. Journal of the Royal
492 Statistical Society: Series B (Methodological), 39(1), 1-22.
- 493 [24] L. A. Shepp and Y. Vardi (1982), Maximum Likelihood reconstruction
494 for emission tomography, IEEE Transactions on Medical Imaging Vol-
495 ume 1, Issue 2, pp. 113-122.
- 496 [25] Y.Vardi, L.A. Shepp, and L. Kaufman (1985), A statistical model for
497 positron emission tomography. Journal of the American statistical As-
498 sociation, 80(389), 8-20.
- 499 [26] K. Lange and R. Carson (1984). EM reconstruction algorithms for emis-
500 sion and transmission tomography. J Comput Assist Tomogr, 8(2),
501 pp.306-16.
- 502 [27] J. Friedman, T. Hastie, and R. Tibshirani (2001). The elements of sta-
503 tistical learning (Vol. 1, No. 10). New York: Springer series in statistics.
- 504 [28] K.J. Coakley (1991). A cross-validation procedure for stopping the EM
505 algorithm and deconvolution of neutron depth profiling spectra. IEEE
506 Transactions on Nuclear Science, 38(1), 9-15.
- 507 [29] K.J. Coakley and J. Llacer (1991). Use of cross-validation as a stopping
508 rule in emission tomography image reconstruction. In Medical Imaging
509 V: Image Physics (Vol. 1443, pp. 226-233). International Society for
510 Optics and Photonics.
- 511 [30] J. Llacer, E. Veklerov, K.J. Coakley, E.J. Hoffman and J. Nunez (1993)
512 Statistical analysis of maximum likelihood estimator images of human
513 brain FDG PET studies, In IEEE Transactions on Medical Imaging, vol.
514 12, no. 2, pp. 215-231, June 1993
- 515 [31] R.E. Carson, Y. Yan, B. Chodkowski, T. K. Yap, and M. E. Daube-
516 Witherspoon (1994). Precision and accuracy of regional radioactivity
517 quantitation using the maximum likelihood EM reconstruction algo-
518 rithm. IEEE transactions on medical imaging, 13(3), 526-537.

- 519 [32] W. Xia, S.J. Glick, T.S. Pan, E.J. Soares, and D.S. Luo (1997). Region
520 of interest evaluation of SPECT image reconstruction methods using a
521 realistic brain phantom. *IEEE Transactions on Nuclear Science*, 44(3),
522 1336-1341.
- 523 [33] F. Meisenkothen, E.B. Steel, T.J. Prosa, K.T. Henry and R.P. Kolli
524 (2015) Effects of detector dead-time on quantitative analyses involving
525 boron and multi-hit detection events in atom probe tomography, *Ultra-*
526 *microscopy* 159 (Part 1) pp. 101111.
- 527 [34] J.L. Wiza (1979), Microchannel plate detectors, *Nucl. Instrument.*
528 *Methods* 162
- 529 [35] K. Oba, P. Rehak (1981) Studies of high-gain micro-channel plate pho-
530 tomultipliers, *IEEE Trans. Nucl. Sci.* NS-28 pp. 683688.
- 531 [36] M. Liptak, W.G. Sandie, E.G. Shelley, D.A. Simpson (1984) Microchan-
532 nel plate electron multiplier for mass spectrometer applications *IEEE*
533 *Trans Nucl Sci*, NS-31, pp. 780-785
- 534 [37] D.M. Murphy and K. Mauersberger (1985) Operation of a microchannel
535 plate counting system in a mass spectrometer, *Rev. Sci. Instrument.* 56
536 (1985) 220226.
- 537 [38] M. Muller, T. Mikoviny, A. Wisthaler (2014) Detector aging induced
538 mass discrimination and non-linearity effects in PTR-ToF-MS, *Int. J.*
539 *Mass Spectrom.* 365-366 9397.
- 540 [39] A. Mood, F.A. Graybill and D.C. Boes (1974).A Introduction to the
541 theory of statistics, McGraw Hill.
- 542 [40] G. Last and M. Penrose, M. (2017). Lectures on the Poisson process
543 (Vol. 7). Cambridge University Press.
- 544 [41] A.Sen, M. Bachhav, F. Vurpillot, J.M. Mann, P.K. Morgan, T.A. Prus-
545 nick, and J.P. Wharry (2021) Influence of field conditions on quantitative
546 analysis of single crystal thorium dioxide by atom probe tomography.
547 *Ultramicroscopy*, 220, p.113167.
- 548 [42] J.A. Fessler and A.O. Hero (1994). Space-alternating generalized
549 expectation-maximization algorithm. *IEEE Transactions on signal pro-*
550 *cessing*, 42(10), pp.2664-2677.