Innovations in Maximum Likelihood Quantum State Tomography

Theory: S. Glancy, E. Knill

Experiment: T. Gerrits, T. Clement, B. Calkins, A. Lita, A. Miller, A. Migdall, S. W. Nam, and R. Mirin

National Institute of Standards and Technology, USA

Experiences with Maximum Likelihood Quantum State Tomography

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Struggles with Maximum Likelihood Quantum State Tomography

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Plan

- "Schrödinger cat" making experiment
- State of quantum state tomography
- Our work
 - stopping criterion
 - improved maximum likelihood algorithm
 - approximate confidence intervals
- Preliminary cat state data

"Schrödinger Cat" States

- (I'm talking about the) state of a single harmonic oscillator.
- superposition of two coherent (or "displaced vacuum") states

$$|\pm\rangle = \frac{1}{\sqrt{2\pm 2e^{-2|\alpha|^2}}} (|-\alpha\rangle \pm |\alpha\rangle)$$

- |+> has only even numbers of photons.
- $|-\rangle$ has only odd numbers.
- (+| -⟩=0

- This type of Schrödinger cat states have been made in a light field trapped in a cavity, microwaves in a superconducting resonator, motion of a trapped ion, traveling light wave (others?)
- With photon (or phonon) numbers < 10.

"Schrödinger Cat" States

- (I'm talking about the) state of a single harmonic oscillator.
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- This type of Schrödinger cat states have been made in a light field trapped in a cavity, microwaves in a superconducting resonator, motion of a trapped ion, traveling light wave (others?)
- With mean photon (or phonon) numbers < 10.

How to Make Cat States

• Original cat making scheme:

- Use Kerr effect Hamiltonian

$$|\alpha\rangle \longrightarrow \chi \hat{n}^2$$

 $t = \frac{\pi}{2\chi} |-\alpha\rangle + i|\alpha\rangle$

- Current materials have too much absorption and too small χ , but there is hope for EIT methods.
- We need to make cats with a specific optical mode shape, and Kerr effect interactions will disturb the mode.

Yurke and Stoler PRL **57**, 13

Lower Order Nonlinearity + Post Selection















Glancy and Vasconcelos arXiv:0705.2045 8

Photon Subtraction

 $\hat{S}(z)$

click

- Make squeezed light by degenerate down conversion. $\omega_{pump} \rightarrow 2\omega_{squeezed}$
- Send through beam splitter.
- Trigger on observing a photon.
- Works like heralded single photon source, but with stronger squeezing ~3dB.



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perfect cat state $|\alpha|^2 = 0.8$

click

Our Photon Subtraction



- Using superconducting transition edge sensor (TES) photon number resolving detectors.
 - efficiency $\sim 90\%$
 - dark counts limited by black-body radiation
- Subtracting more photons makes a higher fidelity, larger cat, using less squeezing.
- Four Data Sets:
 - n=1 by avalanche photo diode (APD)
 - -n=2 by APD
 - -n=2 by TES
 - -n=3 by TES

Measure by Homodyne Detection



- Vary local oscillator phase \u00f6, observe x.
- Record *N* pairs: $\{(x_m, \phi_m) | m = 1...N\}$.
- Calibrate system efficiency $\eta = \eta_{opt} \eta_{pd} \eta_{mm} \eta_{dc}$
 - η_{opt} =optical components 94.0% ± 0.5%
 - η_{pd} =photo-diodes 97.6% ± 2.2%
 - $\eta_{\rm mm}$ =mode-mismatch 95.0% ± 0.5%
 - $\eta_{\rm dc}$ =dark current 97.9% ± 0.1%
- η ~ 85% ± 3%

Lvovsky & Raymer quant-ph/0511044

Forward Measurement Model

• Relate measurement probabilities to quantum state ρ : $P(x|\phi) = \operatorname{Tr}[\Pi(x,\phi)\rho]$



- $\Pi(x,\phi)$ is an element of a continuous POVM. $\Pi(x,\theta) = \sum_{n} E_{n}(\eta) e^{-i\phi a^{\dagger}a} |x\rangle \langle x| e^{+i\phi a^{\dagger}a} E_{n}(\eta)^{\dagger}$
 - $|x\rangle$ is the harmonic oscillator position eigenstate in photon number basis.
 - $-e^{-i\phi a^{\dagger}a}$ is the phase evolution operator.
 - $E_n(\eta)$ are the Kraus operators for photon loss.

Do Quantum State Tomography

- Using set of observations $\{\Pi_m = \Pi(x_m, \phi_m) | m=1...N\}$, and $P(x|\phi) = \text{Tr}[\Pi(x, \phi)\rho]$
- infer state ρ .
- Choose tomography school:
 - linear inversion
 - maximum likelihood
 - Bayesian inference
 - maximum entropy

Paris & Řeháček (editors) Quantum State Estimation

Maximum Likelihood

• Likelihood function:

$$L(\rho) = \prod_{m=1}^{N} \operatorname{Tr}(\Pi_{m} \rho)^{n_{m}}$$

• Loglikelihood function:

$$\mathcal{L}(\rho) = \sum_{m=1}^{N} n_m \log(\mathrm{Tr}(\Pi_m \rho))$$

- Maximize $\mathcal{L}(\rho)$ to find ρ .
- Respect ρ 's constraints: hermitian, $Tr(\rho)=1$.
- *L* is convex.

RpR Maximum Likelihood

- Iterative scheme:
 - begin with $\rho_0 = \mathcal{N}(I)$ = maximally mixed state
 - at each step *i*, compute

$$R(\rho_i) = \sum_{m=1}^{N} \frac{\Pi_m}{\mathrm{Tr}(\Pi_m \rho_i)}$$

- find next $\rho_{i+1} = \mathcal{N}(R_i \rho_i R_i)$
- at maximum likelihood point $\rho_{\rm ML} = \mathcal{N}(R_{\rm ML}\rho_{\rm ML}R_{\rm ML})$
- *R* is positive and hermitian, so each ρ_i is also hermitian and can be normalized to have trace 1.
- The "diluted algorithm", in which $R \rightarrow I + \epsilon R$, will increase \mathcal{L} if, and ϵ is small enough.
- In practice $\varepsilon \rightarrow \infty$.

Řeháček et al. quant-ph/0611244

R_{\rho}R Virtues

- Always returns a density matrix.
- Has a clear method to incorporate measurement noise by adapting Π_m 's.
- No need to parameterize ρ or use constraint equations.
- Simple implementation.

$R\rho R$ Desiderata

- Stopping criterion
- Faster convergence
- Confidence region for ρ
 - or confidence intervals for observables of ρ .

$R\rho R$ Stopping Criterion

- Just do "a lot" of iterations?
- Compare likelihood found at each iteration?
- Compare fidelity (or trace distance?) between states at each iteration?
- We would like to bound the maximum likelihood using our knowledge of the current ρ .

RpR Stopping Criterion

- Consider subsequent ρ 's: $\rho' = \rho + \varepsilon(\sigma - \rho)$
- Expand \mathcal{L} to first order in \mathcal{E} :

$$\mathcal{L}(\rho') \approx \mathcal{L}(\rho) + \varepsilon \frac{\partial}{\partial \varepsilon} \mathcal{L}(\rho') \bigg|_{\varepsilon=0}$$
$$\mathcal{L}(\rho') \approx \mathcal{L}(\rho) + \varepsilon \operatorname{Tr}(R\sigma) - \varepsilon \operatorname{Tr}(R\rho)$$
$$\mathcal{L}(\rho') \approx \mathcal{L}(\rho) + \varepsilon \operatorname{Tr}(R\sigma) - \varepsilon N$$

• What σ will maximize $\mathcal{L}(\rho')$? $\sigma = |\psi\rangle\langle\psi|$, where $|\psi\rangle$ is the eigenvector of R with the largest eigenvalue.

 $\mathcal{L}(\rho') \approx \mathcal{L}(\rho) + \varepsilon[\max(\operatorname{eig}(R)) - N]$

RpR Stopping Criterion



 $\mathcal{L}(\rho') \approx \mathcal{L}(\rho) + \varepsilon [\max(\operatorname{eig}(R(\rho))) - N]$ $\mathcal{L}(\rho') \approx \mathcal{L}(\rho) + \varepsilon r$ $-r = \max(\operatorname{eig}(R(\rho))) - N$

- \mathcal{L} is convex, so $\mathcal{L}(\rho) + r \ge \mathcal{L}_{\max}$
- \therefore stop iterations when r is small.

R ρ *R* Stopping Criterion $r \ge \mathcal{L}_{max} - \mathcal{L}(\rho)$



≤10 photons, 40,000 measurements

- Bound is not very tight.
- Bounding $\mathcal{L}_{\rm ML}$ is good, but I wish we had a bound on the difference between ρ_i and $\rho_{\rm ML}$.

- Can we find an algorithm that converges faster?
- Strategy:
 - use traditional ideas of gradient ascent,
 - trust region / quadratic approximation of \mathcal{L} ,
 - over-parameterize ρ to make optimization unconstrained.
 - To stay within trust region, restrict step size of each iteration.

Parameterization of *ρ*:

$$\rho_{i+1} = \mathcal{N}\left(\left(\rho_{i}^{1/2} + A\right)\left(\rho_{i}^{1/2} + A^{\dagger}\right)\right)$$

- *A* may be any matrix.
- Ensures ρ_{i+1} is a density matrix.
- Increases parameter space from d²-1 to 2d², where d is Hilbert space dimension.

- Quadratic approximation of *L*:
 - $-\rho_{i+1} = \rho_i + \Delta$, where Δ is 2nd order in *A*.

$$\mathcal{L}_{Q}(\rho_{i+1}) \approx \mathcal{L}(\rho_{i}) + \operatorname{Tr}(R_{i}\Delta) - \frac{1}{2} \sum_{m=1}^{N} n_{m} \left(\frac{\operatorname{Tr}(\Pi_{m}\Delta)}{\operatorname{Tr}(\Pi_{m}\rho_{i})} \right)^{2}$$

- Write A as a $2d^2$ element real vector \vec{A} . $\mathcal{L}_Q(\rho_{i+1}) \approx \mathcal{L}(\rho_i) + \vec{v}^T \vec{A} + \frac{1}{2} \vec{A}^T M \vec{A}$
- Choose maximum step size: $s = \vec{A}^T \vec{A}$
- Maximize $\mathcal{L}_Q(\rho_{i+1})$ subject to constraint $s \ge A^T \overline{A}$:

$$\vec{A}(\lambda) = (2\lambda I - M)^{-1} \vec{v}$$

• λ is a Lagrange multiplier, which we set to $\lambda = \max(\operatorname{eig}(M))$ and increase if necessary.

- 1. Choose step size *s*=1.
- 2. From ρ_i , calculate v, M.
- 3. $\lambda = \max(\operatorname{eig}(M)), \ \vec{A}(\lambda) = (2\lambda I M)^{-1} \vec{v}.$
- 4. Check step size: if $\vec{A}(\lambda)^T \vec{A}(\lambda) \ge s$, increase lambda and goto 3.
- 5. Calculate new $\rho_{i+1} = \mathcal{N}((\rho_i^{1/2} + A)(\rho_i^{1/2} + A^{\dagger})).$
- 6. If (exact) $\mathcal{L}(\rho_{i+1}) \leq \mathcal{L}(\rho_i)$, reduce *s* and goto 4.

RGA vs. RpR Competition



20 photons, 2,000 measurements



$= \frac{10^{5}}{10^{0}} \frac{17 \text{ RpR/RGA}}{10^{-5}}$

10 photons, 20,000 measurements

20 photons, 20,000 measurements



- If stopping r is small enough, RGA is faster.
- For high dimensions $R\rho R$ can be faster for larger r.

RGA & RpR Cooperation

 Use RρR for time equal to one RGA iteration, then switch to RGA.



20 photons, 2,000 measurements





20 photons, 2,0000 measurements



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- 1. Use experimental data $\{(x_m, \phi_m) | m=1...N\}$ to find ρ_{ML} .
- 2. Use ρ_{ML} to simulate *B* new data sets, each of which uses the same $\{\phi_m | m=1...N\}$. For each ϕ_m , sample from $P(x | \phi_m) = \text{Tr}[\Pi(x, \phi_m) \rho_{\text{ML}}]$
- 3. For each simulated data set, infer $\{\rho_{ML}^{(b)}|b=1...B\}$.
- 4. Use $\{\rho_{ML}^{(b)}|b=1...B\}$ to calculate parameter of interest $F(\rho)$.
- 5. Obtain distribution of $F^{(b)}=F(\rho_{ML}^{(b)})$.

Resampling for fidelity with ideal cat state:



subtracting 1 photon

- 324,510 measurements
- B=100 data sets
- Long red line is maximum likelihood.

Shorter red lines mark central 68 percentile.

- Resampling is biased toward lower fidelity with pure state.
- We also see bias toward less negative Wigner function values.

• Resampling for fidelity with ideal cat state:



subtracting 2 photons

41,223 measurements

B=100 data sets

Long red line is maximum likelihood.

Shorter red lines mark central 68 percentile.

- Resampling is usually biased toward lower fidelity with pure state.
- We also see bias toward less negative Wigner function values.

• Resampling for fidelity with ideal cat state:



subtracting 3 photons 1087 measurements B = 100 data sets Central red line is

maximum likelihood.

Outer red lines mark central 68 percentile.

- Resampling is biased toward lower fidelity with pure state.
- We also see bias toward less negative Wigner function values.

• Resampling for fidelity with ideal cat state:



subtracting 3 photons
1087 measurements
B = 800 data sets
Central red line is maximum likelihood.
Outer red lines mark central 68 percentile.

- Resampling is biased toward lower fidelity with pure state.
- We also see bias toward less negative Wigner function values.

- Can we correct for the bias?
- Given: $F(\rho_{ML}) = F_{ML}, P(F|\rho_{ML}), F_{ML}^{(1)}, F_{ML}^{(u)}$



Efron Canadian J. Statistics **9**, 139 J. Amer. Statist. Assoc. **82**, 171

- Can we correct for the bias?
- Given: $F(\rho_{ML}) = F_{ML}, P(F|\rho_{ML}), F_{ML}^{(1)}, F_{ML}^{(u)}$
- Hypothesize ρ_0 , a candidate for the true state ρ_T .
- Imagine $F(\rho_0) = F_0, P(F|\rho_0), F_0^{(1)}, F_0^{(u)}$.



- Can we correct for the bias?
- Given: $F(\rho_{ML}) = F_{ML}, P(F|\rho_{ML}), F_{ML}^{(1)}, F_{ML}^{(u)}$
- Hypothesize ρ_0 , a candidate for the true state ρ_T .
- Imagine $F(\rho_0) = F_0, P(F|\rho_0), F_0^{(1)}, F_0^{(u)}$.
- Assume $P(F|\rho_0) = P(F-f_0|\rho_{ML})$.



• Resampling for fidelity with ideal cat state:



subtracting 3 photons
1087 measurements
B = 800 data sets
Central red line is maximum likelihood.
Outer red lines mark central 68 percentile.

- Resampling is biased toward lower fidelity with pure state.
- We also see bias toward less negative Wigner function values.

• Resampling from state "close" to maximum likelihood



subtracting 3 photons

1087 measurements

B = 800 data sets

Central red line is fidelity of state used to generate data.

Outer red lines mark central 68 percentile.

• Histograms look similar, but clearly $P(F|\rho_0) = P(F-f_0|\rho_{ML})$ is not exactly true.

- Can we correct for the bias?
- Given: $F(\rho_{ML}) = F_{ML}, P(F|\rho_{ML}), F_{ML}^{(1)}, F_{ML}^{(u)}$
- Hypothesize ρ_0 , a candidate for the true state ρ_T .
- Imagine $F(\rho_0) = F_0, P(F|\rho_0), F_0^{(1)}, F_0^{(u)}$.
- Assume $P(F|\rho_0) = P(F-f_0|\rho_{ML})$.



- Can we correct for the bias?
- Given: $F(\rho_{ML}) = F_{ML}, P(F|\rho_{ML}), F_{ML}^{(1)}, F_{ML}^{(u)}$
- Hypothesize ρ_0 , a candidate for the true state ρ_T .
- Imagine $F(\rho_0) = F_0, P(F|\rho_0), F_0^{(1)}, F_0^{(u)}$.
- Assume $P(F|\rho_0) = P(F-f_0|\rho_{ML})$.
- If ρ_o is a good hypothesis, $F_o^{(l)} < F_{ML} < F_o^{(u)}$



- What are the allowed locations for F_{o} , such that $F_{o}^{(l)} < F_{ML} < F_{o}^{(u)}$?
- $F_{\rm ML} + F_{\rm o} F_{\rm o}^{(u)} < F_{\rm o}$

 $F_{ML}^{(1)}$ $F_{O}F_{MLO}F_{MLO}^{(1)}$

P(F)



- What are the allowed locations for F_{o} , such that $F_{o}^{(l)} < F_{ML} < F_{o}^{(u)}$?
 - $F_{\rm ML} + F_{\rm o} F_{\rm o}^{(u)} < F_{\rm o} < F_{\rm ML} + F_{\rm o} F_{\rm o}^{(1)}$ P(F) $I_{\mathrm{ML}}^{F(1)}F_{\mathrm{ML}}^{(1)}$ $F_{0}^{(u)}$ $F_{\rm ML}^{(l)}$

- What are the allowed locations for F_{o} , such that $F_{o}^{(l)} < F_{ML} < F_{o}^{(u)}$?
- $F_{\rm ML} + F_{\rm o} F_{\rm o}^{(\rm u)} < F_{\rm o} < F_{\rm ML} + F_{\rm o} F_{\rm o}^{(\rm l)}$
- Because $P(F|\rho_0) = P(F-f_0|\rho_{ML})$,
 - $F_{o} F_{o}^{(u)} = F_{ML} F_{ML}^{(u)}$
 - $F_{o} F_{o}^{(1)} = F_{ML} F_{ML}^{(1)}$
- : $2F_{\rm ML} F_{\rm ML}^{(u)} < F_{\rm o} < 2F_{\rm ML} F_{\rm ML}^{(l)}$



- What are the allowed locations for F_{o} , such that $F_{o}^{(l)} < F_{ML} < F_{o}^{(u)}$?
- $F_{\rm ML} + F_{\rm o} F_{\rm o}^{(\rm u)} < F_{\rm o} < F_{\rm ML} + F_{\rm o} F_{\rm o}^{(\rm l)}$
- Because $P(F|\rho_0) = P(F-f_0|\rho_{\rm ML})$,
 - $F_{o} F_{o}^{(u)} = F_{ML} F_{ML}^{(u)}$
 - $F_{o} F_{o}^{(1)} = F_{ML} F_{ML}^{(1)}$
- : $2F_{\rm ML} F_{\rm ML}^{(u)} < F_{\rm o} < 2F_{\rm ML} F_{\rm ML}^{(l)}$
- $F_{\rm T}^{(1)} = 2F_{\rm ML} F_{\rm ML}^{(u)}, F_{\rm T}^{(u)} = 2F_{\rm ML} F_{\rm ML}^{(1)}$



Bias Correcting Our Data



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Bias Correcting Our Data



- We must also include calibration uncertainty $\eta \sim 85\% \pm 3\%$.
- Choosing $\eta = 82\%$ or 88% shifts F_{ML} by ~1.5%.
- So, I have increased the size of the data squares.

Conclusions

- MaxLikelihood stopping criterion:
 - $r = \max(\operatorname{eig}(R(\rho))) N$
 - bounds likelihood: $\mathcal{L}_{\max} \leq \mathcal{L}(\rho) + r$
- Regularized Gradient Ascent maximization algorithm.
 - faster convergence, but may not be practical in all cases
 - can optimize any convex function of ρ .
- Parametric bootstrap resampling with bias correction
 - correct low-purity bias of MaxLikelihood inference.
 - requires strong assumptions.
- Created approximate cat states by subtracting 3 photons.
 - $\langle n \rangle$ is fairly large, but fidelity needs improvement
 - requires higher purity squeezing

Open Problems

- More rigorous confidence intervals that don't require resampling.
 - Ask me about what we have tried that didn't work, and
 - ideas we have for achieving this.
- Test for number of photons required in density matrix
 - too many photons may cause "over fitting" problems.
- Fast method for Bayesian inference of ρ .
- How to make high purity, single mode, pulsed, squeezed light.

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